



Statistical Theory of Extremes

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Part 1

Probabilistic Patterns of Univariate Statistical Extremes

Annex 1

On The “Duality” between Extremes and Sums

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For simplicity we will deal with some “duality” between sums (or averages) and maxima, the translation to minima being obvious from the relation

$$\min_{1 \leq i \leq n} \{X_i\} = -\max_{1 \leq i \leq n} \{-X_i\}.$$

The “duality” is expressed by the two columns in correspondence, where there are various gaps. $F(\cdot)$, $F(\cdot, \cdot)$, ... and $\varphi(\cdot)$, $\varphi(\cdot, \cdot)$, ... will denote the distribution functions and the characteristic functions.

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Sums	Maxima
$S_k = \sum_{i=1}^k X_i$ $\varphi_X(t) = M_X(e^{itX}): \text{ch. f. of } X$ $X_i \text{ indep.: } \varphi_{S_k}(t) = \prod_{i=1}^k \varphi_i(t)$ $X_i \text{ i. i. d.: } \varphi_{S_k}(t) = \varphi^k(t)$ $(X, Y) \text{ indep.: } \varphi_{aX+bY}(t) = \varphi_X(at) \varphi_Y(bt)$ $(+, \cdot)$ $M(aX+b) = a M(X) + b$ $V(aX+b) = a^2 V(X)$ If $\{X_i\}$ i.i.d. have μ, σ^2 then $\frac{\varphi(t) - e^{it\mu}}{(s_k - k\mu)/\sqrt{k}\sigma} \rightarrow (e^{-t^2/2})$ (Central Limit Theorem); in the general case “sometimes” the ch. f. of the normal law may be substituted by that of an indefinitely divisible law ; $\mu = \varphi'_X(0)/i, \sigma^2 = \varphi'_X(0)^2 - \varphi''_X(0)$, in the usual case. If (X, Y) has a binormal distribution standard normal margins and correlation coefficient then $Z = aX + bY$ has a normal distribution $N(x/\sigma(a, b))$ with $\sigma(a, b) = a^2 + b^2 + 2\rho(a, b)$. If $\rho = 0$ (independence), $\sigma(a, b) = 1$ iff $a^2 + b^2 = 1$; in the case $C(X, Z) = a$.	$M_k = \max_{1 \leq i \leq k} \{X_i\}$ $F_X(t) = \text{Prob}\{X \leq x\}: \text{d. f. of } X$ $X_i \text{ indep.: } F_{M_k}(x) = \prod_{i=1}^n F_i(x)$ $X_i \text{ i. i. d.: } F_{M_k}(x) = F^k(x)$ $(X, Y) \text{ indep.: } F_{\max(X+a, Y+b)} = F_X(x-a)F_Y(y-b)$ $(\max, +)$ If $\{X_i\}$ are i.i.d. there “sometimes” exist $(\lambda_k, \delta_k > 0)$ such that $\text{Prob}\{(M_k - \lambda_n)/\delta_n \leq x\} = F^k(\lambda_k + \delta_k x) \rightarrow \tilde{L}(x), \tilde{L}(x)$ then being $\Psi_\alpha(x), \Lambda(x)$ or $\Phi_\alpha(x)$; for $\Lambda(x)$ we have $n(1 - F(\lambda_k)) \rightarrow 1, k(1 - F(\lambda_k + \delta_k)) \rightarrow e^{-1}$ or $\delta_n \sim 1/n F'(\lambda_n)$; there are corresponding results for Φ_α and Φ_α. If (X, Y) has a bivariate distribution with reduced Gumbel margins then $Z = \max(X - a, Y - b)$ has a Gumbel distribution $\Lambda(z - \lambda(a, b))$ with $\lambda(a, b) = \log\{(e^{-a} + e^{-b})k(b - a)\}$ If $k(w) = 1$ (independence), $\lambda(a, b) = 0$ if $e^{-a} + e^{-b} = 1$; in that case $\text{Prob}\{Z \leq X - a\} = \text{Prob}\{X - b \leq X - a\} = e^{-a}$.
