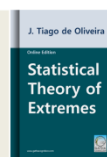




# Statistical Theory of Extremes

Homepage: <http://www.gathacognition.com/book/gcb14>  
<http://dx.doi.org/10.21523/gcb14>



## Part 4

### Multivariate Extremes

## Exercises

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| Editor(s)              | Published Online |
|------------------------|------------------|
| J.C. Tiago de Oliveira | 23 June 2017     |

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- 4.1. Consider the EMS sequence  $\{Z_k\}$ . Obtain the expression of  $\text{Prob}\{\min(Z_1, \dots, Z_k) \geq z\} = Q_k(z, z|\theta)$

where

$$Q_k(x, y|\theta) = \text{Prob}\left\{\min_{1 \leq i \leq k-1} (Z_i, \dots, Z_{k-1}) \geq x, Z_k \geq y\right\}$$

$$= Q_{k-1}(x, \max(x, y - \log \theta)) \wedge (y - \log(1 - \theta)).$$

Obtain its expression when  $x < y - \log \theta$  and  $x > y - \log \theta$ .

- 4.2. Consider the general EMS sequence  $X_k = \lambda + \delta Z_k$  and the seek estimators of  $(\lambda, \delta, \theta)$ . We have  $\min_{1 \leq i \leq k} (X_i - X_{i-1}) \xrightarrow{P} \delta \log \theta$  which gives an (over-) estimator of  $\Delta = \delta \log \theta$ . Using this result obtain, supposing  $\Delta^* = \Delta$ , an estimator of  $(\lambda, \delta)$ .
- 4.3. Translate maxima results to minima results and vice-versa; use, in particular, the relation between the distribution function and the survival function.

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Originally published in 'Statistical Analysis of Extremes', 1997, 2016

<http://dx.doi.org/10.21523/gcb1.17026>

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4.4. Study the processes of oldest ages of death for men and women in Sweden (Table.1 and Tables 1 + 2) see [Exercise of Part 3](#), trying to fit any of the random sequences that may have trend and/or oscillations. In particular they may be considered as sliding processes of maxima, with Gumbel margins.

4.5. Consider a reduced extremal process  $Z(t), t > 0$  observed at (non-random) instants  $(0 <) t_1 < t_2 < \dots < t_n < \dots$ . The best (least squares) predictor of  $Z(t_{n+1})$  ( $t_n < t_{n+1}$ ) of the form  $Z_1^*(t) = Z(t_n) + a$  has  $a = \log(t_{n+1}/t_n)$ . The best (i.e., least squares) linear predictor of  $Z(t_{n+1})$  based only on  $Z(t_n)$  is of the form  $Z^{**}(t_{n+1}) = \alpha + \beta Z(t_n)$ ; the minimization of  $MSE = M(Z^{**}(t_{n+1}) - Z(t_{n+1}))^2$  leads to

$$Z^{**}(t_{n+1}) = \gamma + \log t_{n+1} + \rho(t_n/t_{n+1}) (Z(t_n) - \gamma + \log t_n).$$

They are both unbiased (i.e., with the same mean value) and

$$MSE(Z^*(t_{n+1})) = \frac{\pi^2}{3} (1 - \rho(t_n/t_{n+1})) \text{ and } MSE(Z^{**}(t_{n+1})) = \frac{\pi^2}{6} (1 - \rho^2(t_n/t_{n+1})).$$

Note that  $Z^{**}(t_{n+1})$  can be compared with the simpler linear predictor  $Z^*(t_{n+1}) = Z(t_n) + \log \frac{t_{n+1}}{t_n}$ . The efficiency is

$$\frac{MSE(Z^*(t_{n+1}))}{MSE(Z^{**}(t_{n+1}))} = \frac{1 + \rho(t_n/t_{n+1})}{2} < 1.$$

None of the predictors is invariant for linear transformations, and so they can't be used to predict the general extremal processes  $X(t) = \lambda + \delta Z(t)$ .

4.6. Consider the extremal processes  $X(t) = \lambda + \delta Z(t), t \geq 0$ , where  $Z(t)$  is the reduced extremal process. Suppose observations are made at (non-random) instants  $(0 <) t_1 < t_2 < \dots < t_n < \dots$ .

Obtain the expression of the quasi-linear predictor of  $X(t_{n+1})$  when  $X(t_i), i = 1, \dots, n$ , are known. The quasi-linear predictor of  $X(t_{n+1})$  based on the last two observations  $X(t_{n-1})$  and  $X(t_n)$  ( $\geq X(t_{n-1})$ ) is, obviously,

$$X^*(t_{n+1}) = X(t_n) + \beta(X(t_n) - X(t_{n-1})).$$

The best (least squares) predictor of  $X(t_{n+1})$ , minimizing

$$\text{MSE}(X^*(t_{n+1})) = M(X^*(t_{n+1}) - X(t_{n+1}))^2 ,$$

is given by

$$\beta = \frac{M((X(t_{n+1}) - X(t_n))(X(t_n) - X(t_{n-1})))}{M((X(t_n) - X(t_{n-1}))^2)}$$

and

$$\text{MSE}(X^*(t_{n+1})) = M((X(t_{n+1}) - X(t_n))^2) - \beta^2 M((X(t_n) - X(t_{n-1}))^2) ,$$

which can be expressed in mean values and covariance of the process.

- 4.7. Define an extremal process of Weibull minima, using the conversion between the Weibull distribution of minima and the Gumbel distribution of maxima.
- 4.8. Consider a max-compound Poisson stochastic process,  $X(t) = \max\{Y_i\}$ , where  $N(t)$  is a Poisson process with intensity  $v$  and  $Y_0, Y_1, \dots, Y_k, \dots$ , is a sequence of independent random variables such that  $Y_0$  has the distribution function  $G_0(x)$  and  $Y_j (j \geq 1)$  have the distribution function  $G(x)$ . Show that  $\text{Prob}\{X(t) \leq x\} = G_0(x) \exp \{-v t(1 - G(x))\}$  and that for no choice of  $(G_0, G)$  can we have  $\text{Prob}\{X(t) \leq x\} = \Lambda(\alpha(t) + \beta(t)x)$ .
- 4.9. Analyse the same question for max-filtered Poisson processes and max-renewal point processes.

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